

OBSERVE THE METHODS AND MEASURES USED IN TWO LEVEL AND TWO OCCASIONS SEQUENTIAL SAMPLING

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Abstract:

Using two occasions of successive sampling, an effective estimation approach has been established in this study to estimate the current population mean. The current population means is estimated using a method based on exponential regression, and an ideal replacement approach is recommended. The proposed estimator's superiority to sample mean estimators and natural sequential sampling estimators is demonstrated empirically. Findings have been interpreted, and appropriate suggestions have been given.

Keywords: sampling, two-occasion, successivesampling

1. Introduction:

A one-time survey conducted on a single occasion does not provide information about the nature or pattern of change of the characteristic over different occasions, the precise estimates of the characteristic over all occasions, or on the most recent occasion, when the characteristic under study of a finite population changes over time. Instead, it only provides information about the characteristic of the surveyed population for the given occasion. In order to get around this problem, sampling is done repeatedly in order to produce accurate estimates of population parameters at various times. The theory of sequential sampling appears to have been introduced by Jessen (1942), who was a pioneer in utilizing all of the data gathered from earlier studies to improve current estimations. Pattersons (1950), Rao and Graham (1964), Gupta (1979), and Das (1982), among others, have added to this idea. Sen (1971) used data on two auxiliary variables that were easily accessible on a previous occasion to create estimators of the population mean on the occasion in question. Sen expanded his study for a number of auxiliary variables in 1972 and 1973.

In two occasions successive sampling, Singh et al. (1991) and Singh and Singh (2001) employed the supplementary information on present occasion to estimate the current population mean. The theory for h-occasion successive sampling was expanded by Singh (2003). In many circumstances,

information on an auxiliary variable may be easily accessible on both the first and second occasions. For instance, in a transportation survey, the tonnage (or seat capacity) of each vehicle or ship is known. Additional instances may be given where the information on auxiliary variables is accessible on both occasions of two-occasion successive sampling. A variety of estimators of the population mean on current (second) occasions in two occasion successive sampling have been proposed by Feng (1997), Birader (2001), Singh (2005), Singh (2006), Singh (2008, 2010), Singh (2009), Singh (2009), Singh (2010), Singh (2011), and Singh (2013), among others, using the auxiliary information on both occasions.

The goal of the current work is to suggest a more accurate estimator of the current population mean in two-occasion sequential sampling using data on two stable auxiliary variables that are easily accessible on both occasions, as a follow-up to the aforementioned justifications. An exponential regression type estimator of the current population mean in two-occasion successive sampling has been presented using the data on the two auxiliary variables. When there is no matching from the prior instance, the proposed estimator's properties are investigated, and relative efficiency comparisons are done with the sample mean estimator and the natural sequential sampling estimator when no auxiliary information is employed. Empirical research has been done, and the results demonstrate the proposed estimator's performance has improved significantly. The results have been expertly interpreted, and pertinent suggestions have been given.

Properties of the Wished-for Estimator

Partiality and Unkind Square Error

Meanwhile the estimator's T_u and T_m remain exponential besides exponential deterioration kind estimators, they remain biased estimators of the population mean \bar{Y} . Then, the resulting estimator T is too a biased estimator of \bar{Y} . The prejudice $B(\cdot)$ and mean right-angled error $M(\cdot)$ of the estimator T is resulting up to the principal order of guesses under large sample supposition and exposed in the next theorems:

Theorem 1. *Partiality of the estimator T to the principal order of calculations is obtained as*

$$(T) = \varphi(T_u) + (1 - \varphi)B(T_m)$$

Where

$$B(T_u) = \left(\frac{1}{u} - \frac{1}{N} \right) \left\{ \alpha_{0100} - \frac{1}{2} \bar{Y} \left(\frac{\alpha_{0010}}{Z_1} + \frac{\alpha_{0001}}{Z_2} \right) - \frac{1}{2} \frac{\alpha_{0010}}{Z_1} - \frac{1}{2} \frac{\alpha_{0110}}{Z_1 Z_2} + \frac{3}{8} \left(\frac{\alpha_{0020}}{Z_1^2} + \frac{\alpha_{0002}}{Z_2^2} \right) \right\}$$

Where

$$\alpha pqr s = E[(x_i - \bar{X})p(y_i - \bar{Y})^q(z_{1i} - \bar{Z}_1)^r(z_{2i} - \bar{Z}_2)^s]; \quad ((p, q, r, s) \geq 0 \text{ are integers}).$$

Theorem 2. Mean square fault of the estimator T to the leading degree of guesses is found as

$$(T) = \varphi^2(T_u) + (1-\varphi)^2 M(T_m) + 2\varphi(1-\varphi)C(T_u, T_m)$$

Lowest mean square blunders of the estimator T

Subsequently the nasty square blunder of the estimator T in calculation is the meaning of the indefinite constant (scalar) φ , so, it is minimized by respect to φ and then the optimum value of φ is found as

$$\varphi_{opt} = \frac{M(T_m) - c(T_u, T_m)}{M(T_u) + M(T_m) - 2c(T_u, T_m)}$$

From overheatdequation, replacing the value of φ_{opt} in calculation we grow the optimum mean four-sided error of the estimator T as

$$M\varphi_{opt} = \frac{M(T_u) + M(T_m) - \{c(T_u, T_m)\}^2}{M(T_u) + M(T_m) - \{2c(T_u, T_m)\}^*}$$

the basic values of φ_{opt} and $M(T)_{opt}$ are got as

$$\begin{aligned}\varphi_{opt} &= \frac{\mu(A_{12} + \mu A_{11})}{A_1 - \mu A_{14} + \mu^2 A_{15}} \\ M\varphi_{opt} &= \frac{(A_{22} - \mu A_{23} + \mu^2 A_{21}) S_y^2}{(A_1 - \mu A_{14} + \mu^2 A_{15}) n}\end{aligned}$$

Optimum Replacement strategy of the estimator T

The finest mean square error $(T)_{pt}$. in reckoning is a meaning μ (fraction of sample to be drawn anew on the next occasion). It is an imperative factor in dropping the rate of the examination, therefore, to fix the optimum cost of μ so that \bar{Y} can be estimated thru maximum meticulousness and tiniest cost, we abate $(T)_{pt}$. with reverence to μ which consequences in a quadratic equation in μ , which is made known as

$$\mu^2 D_1 + 2\mu D_2 + D_3 = 0$$

Table2: Idealvalues of μ_0 in additionPRE'sofTthrurespect to \bar{y}_n and \bar{y}_0 on behalf of f= 0.1

also $P_{z_1 z_2} = 0.5$

p_0	p_1	ρ_{yx}	0.6	0.7	0.8	0.9
0.4	0.4	μ_0	0.3663	0.4658	0.5425	0.6326
		E1	132.78	139.32	149.92	168.88
		E2	118.03	117.19	116.61	115.957
	0.5	μ_0	0.3881	0.4843	0.5535	0.6404
		E1	144.61	152.08	164.33	186.81
		E2	128.54	127.93	127.81	128.26
	0.6	μ_0	0.2343	0.4617	0.5395	0.6258
		E1	160.87	168.63	182.22	207.56
		E2	142.99	141.85	141.73	142.51
	0.8	μ_0	0.6109	0.7156	*	0.4037
		E1	209.73	217.57		268.87
		E2	186.43	183.02		184.61
0.5	0.4	μ_0	0.3881	0.4843	0.5535	0.6404
		E1	144.61	152.08	164.33	186.81
		E2	128.54	127.93	127.81	128.26
	0.5	μ_0	0.5134	0.5382	0.5782	0.6512
		E1	158.43	167.02	181.04	207.33
		E2	140.83	140.49	140.81	142.35
	0.6	μ_0	*	0.6170	0.5921	0.6440
		E1		187.14	202.46	231.69
		E2		157.42	157.46	159.08
	0.8	μ_0	0.4351	0.4520	0.4815	0.5350
		E1	243.59	253.93	272.12	305.72
		E2	216.53	213.60	211.65	209.91
0.6	0.4	μ_0	0.2343	0.4617	0.5395	0.6258
		E1	160.87	168.63	182.22	207.56
		E2	142.99	141.85	141.73	142.51
	0.5	μ_0	*	0.6170	0.5921	0.6440
		E1		187.14	202.46	231.69
		E2		157.42	157.46	159.08
	0.6	μ_0	0.2284	*	0.6820	0.6528
		E1	199.23		230.76	262.27
		E2	177.09		179.48	180.08
	0.8	μ_0	0.3284	0.3282	0.3084	*
		E1	293.80	304.99	324.79	
		E2	261.16	256.56	252.61	
0.7	0.4	μ_0	0.7220	0.0152	0.4638	0.5773
		E1	180.09	193.05	205.64	232.83
		E2	160.08	162.39	159.94	159.86
	0.5	μ_0	0.3863	0.1655	0.6163	0.6176
		E1	203.59	212.35	231.19	262.44
		E2	180.97	178.63	179.82	180.19
	0.6	μ_0	0.3008	0.2300	*	0.6855
		E1	234.84	244.03		305.28
		E2	208.75	205.28		209.60
	0.8	μ_0	0.2521	0.2527	0.2479	0.1987
		E1	389.79	402.30	425.68	465.21
		E2	346.48	338.41	331.08	319.41
0.8	0.4	μ_0	0.6109	0.7156	*	0.4037
		E1	209.73	217.57		268.87
		E2	186.43	183.02		184.61
	0.5	μ_0	0.4351	0.4520	0.4815	0.5350
		E1	243.59	253.93	272.12	305.72
		E2	216.53	213.60	211.65	209.91
	0.6	μ_0	0.3284	0.3282	0.3084	*
		E1	293.80	304.99	324.79	
		E2	261.16	256.56	252.61	
	0.8	μ_0	0.1530	0.1543	0.1559	0.1541
		E1	695.75	712.60	747.82	811.09
		E2	618.44	599.43	581.63	556.90

Table3: Optimum standards of μ_0 besides PRE's of T by respect to \bar{y}_n and \bar{y}_o on behalf of $f = 0.1$

besides $P_{z_1 z_2} = 0.7$

p_0	p_1	p_{yx}	0.6	0.7	0.8	0.9
0.4	0.4	μ_0	0.3733	0.4623	0.5416	0.6350
		$E1$	115.46	121.05	129.85	145.10
		$E2$	102.63	101.83	100.99	100.63
	0.5	μ_0	0.4348	0.5036	0.5701	0.6580
		$E1$	121.73	128.20	138.33	156.35
		$E2$	108.20	107.84	107.59	107.35
	0.6	μ_0	0.4754	0.5277	0.5826	0.6630
		$E1$	130.21	137.28	148.40	168.45
		$E2$	115.74	115.48	115.42	115.65
	0.8	μ_0	0.4697	0.5214	0.5597	0.6211
		$E1$	155.73	163.24	175.28	196.54
		$E2$	138.42	137.32	136.33	134.94
0.5	0.4	μ_0	0.4348	0.5036	0.5701	0.6580
		$E1$	121.73	128.20	138.33	156.35
		$E2$	108.20	107.84	107.59	107.35
	0.5	μ_0	0.5402	0.5668	0.6078	0.6798
		$E1$	128.24	135.37	146.51	166.62
		$E2$	113.99	113.87	113.95	114.40
	0.6	μ_0	0.6662	0.6300	0.6383	0.6886
		$E1$	138.03	145.17	156.74	177.80
		$E2$	122.70	122.11	121.90	122.08
	0.8	μ_0	*	*	0.7570	0.7018
		$E1$			187.91	206.83
		$E2$			146.15	142.01
0.6	0.4	μ_0	0.4754	0.5277	0.5826	0.6630
		$E1$	130.21	137.28	148.40	168.45
		$E2$	115.74	115.48	115.42	115.65
	0.5	μ_0	0.6662	0.6300	0.6383	0.6886
		$E1$	138.03	145.17	156.74	177.80
		$E2$	122.70	122.11	121.90	122.08
	0.6	μ_0	0.9917	0.7648	0.7001	0.7089
		$E1$	153.84	158.00	168.61	189.21
		$E2$	136.75	132.91	131.14	129.91
	0.8	μ_0	*	*	*	0.8236
		$E1$				229.67
		$E2$				157.69
0.7	0.4	μ_0	0.4995	0.5369	0.5808	0.6512
		$E1$	141.27	148.70	160.46	181.56
		$E2$	125.57	125.08	124.80	124.66
	0.5	μ_0	0.9592	0.7287	0.6748	0.6915
		$E1$	153.88	158.77	169.82	190.63
		$E2$	136.79	133.55	132.08	130.89
	0.6	μ_0	*	*	0.8054	0.7403
		$E1$			186.82	204.73
		$E2$			145.30	140.56
	0.8	μ_0	*	*	*	*
		$E1$				
		$E2$				
0.8	0.4	μ_0	0.4697	0.5214	0.5597	0.6211
		$E1$	155.73	163.24	175.28	196.54
		$E2$	138.42	137.32	136.33	134.94
	0.5	μ_0	*	*	0.7570	0.7018
		$E1$			187.91	206.83
		$E2$			146.15	142.01
	0.6	μ_0	*	*	*	0.8236
		$E1$				229.67
		$E2$				157.69

Table4: Optimal values of μ_0 besides PRE's of T with the esteem to \bar{y}_n and \bar{y} for f= 0.1 to $P_{z_1 z_2} = 0$

ρ_0	ρ_1	ρ_{yx}	0.6	0.7	0.8	0.9
0.4	0.4	μ_0	0.4149	0.5004	0.5588	0.6367
		E1	150.37	158.10	170.68	193.61
		E2	133.66	132.99	132.75	132.93
	0.5	μ_0	0.1001	0.4902	0.5507	0.6293
		E1	169.52	177.91	192.59	220.15
		E2	150.68	149.66	149.79	151.16
	0.6	μ_0	0.5926	*	0.4787	0.5868
		E1	194.39		221.85	253.68
		E2	172.79		172.55	174.18
	0.8	μ_0	0.5683	0.6080	0.7026	*
		E1	288.88	299.01	316.56	
		E2	256.78	251.52	246.21	
0.5	0.4	μ_0	0.1001	0.4902	0.5507	0.6293
		E1	169.52	177.91	192.59	220.15
		E2	150.68	149.66	149.79	151.16
	0.5	μ_0	0.4858	0.5091	0.5481	0.6218
		E1	193.72	204.03	221.59	255.63
		E2	172.20	171.63	172.34	175.51
	0.6	μ_0	0.4920	0.5481	0.4277	0.5721
		E1	229.55	241.27	262.56	303.30
		E2	204.05	202.95	204.21	208.25
	0.8	μ_0	0.4960	0.5237	0.5850	0.7486
		E1	388.39	402.84	430.69	478.98
		E2	345.24	338.87	334.98	328.87
0.6	0.4	μ_0	0.5926	*	0.4787	0.5868
		E1	194.39		221.85	253.68
		E2	172.79		172.55	174.18
	0.5	μ_0	0.4920	0.5481	0.4277	0.5721
		E1	229.55	241.27	262.56	303.30
		E2	204.05	202.95	204.21	208.25
	0.6	μ_0	0.4692	0.5041	0.6236	0.4493
		E1	284.25	298.27	323.29	378.53
		E2	252.67	250.90	251.45	259.90
	0.8	μ_0	0.4217	0.4414	0.4855	0.5829
		E1	635.42	658.45	707.62	803.82
		E2	564.82	553.88	550.37	551.91
0.7	0.4	μ_0	0.5753	0.6569	**	0.4144
		E1	231.24	241.03		302.56
		E2	205.55	202.75		207.74
	0.5	μ_0	0.5012	0.5419	0.6656	0.1449
		E1	285.88	298.94	321.74	385.86
		E2	254.11	251.47	250.24	264.93
	0.6	μ_0	0.4555	0.4837	0.5477	0.7857
		E1	383.81	400.85	433.31	491.18
		E2	341.17	337.19	337.01	337.24
	0.8	μ_0	*	*	*	*
		E1				
		E2				
0.8	0.4	μ_0	0.5683	0.6080	0.7026	*
		E1	288.88	299.01	316.56	
		E2	256.78	251.52	246.21	
	0.5	μ_0	0.4960	0.5237	0.5850	0.7486
		E1	388.39	402.84	430.69	478.98
		E2	345.24	338.87	334.98	328.87
	0.6	μ_0	0.4217	0.4414	0.4855	0.5829
		E1	635.42	658.45	707.62	803.82
		E2	564.82	553.88	550.37	551.91

For diverse selections of connections $\rho_{yx}, \rho_0, \rho_1, \rho_{z1z2}$ and f , Tables 1-4 contemporary the finest values of μ and

the % qualified competences E_1 and E_2 with deference to \bar{y}_n and \bar{y} individually, where $E_1 = \frac{v(\bar{y}_n)}{M(T)*_{opt}} \times 100$

then $E_2 = \frac{M(y)_{opt}}{M(T)*_{opt}} \times 100$

2. Results

Table 1

- For fixed tenets of ρ_0 and ρ_1 , the principles of μ_0 and E_1 stand increasing whereas the values of E_2 do not appear any confident pattern per the aggregate values of ρ_{yx} . This conduct is in arrangement by consequences which elucidate that additional the value of ρ_{yx} , extra the segment of fresh model required on the current occasion.
- Aimed at immovable values of ρ_1 and ρ_y , the standards of E_1 and E_2 remain increasing though no certain patterns remain seen in the standards of μ_0 by the aggregate values of ρ_0 .
- For immovable morals of ρ_0 and ρ_{yx} morals of E_1 besides E_2 are amassed while no fixed patterns stay seen in the ethics of μ_0 thru the growing values of ρ_1 .
- Smallest value of μ_0 is experimental as 0.0142, which specifies that the part of sample to be traded on the existing case is as little as about 1% of the entire taster size, which leads to huge lessening in rate of the examination, which is a decidedly desirable spectacle.

Table 2

- For fixed ideals of ρ_0 and ρ_1 , the ideals of μ_0 and E_1 are swelling while the ideals of E_2 do not VDU any final pattern per the proliferation in the principles of ρ_{yx} . This behavior is in settlement through the ordinary marks in continual sampling which enlightened that the further fraction of different sample continuously the present-day event is required per the intensification in the ideals ρ_{yx} .
- For stationary tenets of ρ_1 and ρ_{yx} , the tenets of E_1 and E_2 stay increasing while no convinced patterns stay visible in the principles of μ_0 per the collective values of ρ_0 .
- Aimed at fixed standards of ρ_0 and ρ_{yx} the standards of E_1 and E_2 are snowballing while no sure trends are understood in the standards of μ_0 if we proliferation the morals of ρ_1 .
- Minimum worth of μ_0 is create as 0.0292, which specifies that the segment of taster to be swapped on the existing case is as low as around 2 percent of the whole sample size, which hints in lessening of survey rate, such behavior is constantly desired in examinations sampler.

Table 3

- For motionlesschoosing of ρ_0 plus ρ_1 , principles of μ_0 besides E_1 expression the collective pattern whereas the valuesof E_2 do not survey any development when we escalation the tenets of ρ_{yx} .These behavior'sbacking theregularscheme of successive test group that new the value of ρ_{yx} ,new the fraction of differentmockupisrequired onthe contemporaryinstance.
- On behalf ofstationaryranges of ρ_1 and ρ ,tenets of E_1 and E_2 are collective while no fashions are appreciated intheprinciples of μ_0 by way ofthecollectivevalues of ρ_0 .
- Slightest value of μ_0 is perceived as0.3833, which point to that the division to be switched on thepresent-dayincident is as low asapproximately 35ofthe 24% total mockupsize,which pointers in fall ofthe assessmentcost.
- This discovery generates inquisitiveness to inspect the behavior of the planned estimator when the supplementary variables remain independent. Aimed at such state of affairs results remain given in Table 4.

Table 4

- When ancillary variables be present uncorrelated, it drinks been perceived that for fixed ranges of ρ_0 and ρ_1 the tenets of μ_0 and E_1 intensification with the intensification in the values of ρ_{yx} , despite the point that no definite arrangements are perceived in the values of E_2 .

3. Conclusions

From the above interpretations and discussions it has been observed that the use of information on two auxiliary variables on estimation stage is highly rewarding in terms of precision of the proposed estimator. The most important point, we have noticed in the present work is the percent relative efficiencies of the proposed estimator are decreasing with the increase in the values of correlation coefficient between auxiliary variables z_1 and z_2 . This phenomenon suggests that if information on more number of mutually least correlated auxiliary variables is used at the estimation stage, more reliable estimates of population parameters may be generated. Looking on the nice behavior of the proposed estimator the survey statisticians may be recommended for its practical applications in their real life problems.

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